

ACSC/STAT 4703, Actuarial Models II

Fall 2015

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Homework Sheet 4
Model Solutions

Basic Questions

1. An insurance company models number of claims an individual makes in a year as following a negative binomial distribution with $\beta = 1.4$, and R an unknown parameter with prior distribution a gamma distribution with $\alpha = 3$ and $\theta = 0.04$.

(a) What is the probability that a random individual makes exactly 3 claims?

If $R = r$, then the conditional probability of making exactly 3 claims is $\frac{r(r+1)(r+2)}{6} \left(\frac{1}{2.4}\right)^r \left(\frac{1.4}{2.4}\right)^3 = 0.03308256 \frac{r(r+1)(r+2)}{2.4^r} = 0.03308256r(r+1)(r+2)e^{-\log(2.4)r}$.

The probability is the expected value of the conditional probability, which is given by

$$\begin{aligned} & \int_0^\infty \left(\frac{r^2}{0.04^3 \Gamma(3)} e^{-\frac{r}{0.04}} \right) \left(0.03308256r(r+1)(r+2)e^{-\log(2.4)r} \right) dr \\ &= \frac{0.03308256}{0.04^3 \Gamma(3)} \int_0^\infty e^{-25r} r^3(r+1)(r+2)e^{-\log(2.4)r} dr \\ &= 258.4575 \int_0^\infty e^{-(25+\log(2.4))r} (r^5 + 3r^4 + 2r^3) dr \\ &= 258.4575 \left(\int_0^\infty r^5 e^{-(25+\log(2.4))r} dr + 3 \int_0^\infty r^4 e^{-(25+\log(2.4))r} dr + 2 \int_0^\infty r^3 e^{-(25+\log(2.4))r} dr \right) \\ &= 258.4575 \left((25 + \log(2.4))^{-6} \Gamma(6) + (25 + \log(2.4))^{-5} \Gamma(5) + (25 + \log(2.4))^{-4} \Gamma(4) \right) \\ &= 0.004097388 \end{aligned}$$

(b) The company now observes the following claim frequencies:

Number of claims	Frequency
0	584
1	90
2	36
3	12
4	3
5	3
6	1

What is the probability that $R > 0.4$?

The posterior probability density is

$$\frac{\pi(r)r^{145}(r+1)^{55}(r+2)^{19}(r+3)^7(r+4)^4(r+5)\left(\frac{1}{2.4}\right)^{729r}}{\int_0^\infty \pi(r)r^{145}(r+1)^{55}(r+2)^{19}(r+3)^7(r+4)^4(r+5)\left(\frac{1}{2.4}\right)^{729r} dr}$$

The probability that $R > 0.4$ is therefore

$$\frac{\int_{0.4}^\infty r^2 e^{-25r} r^{145}(r+1)^{55}(r+2)^{19}(r+3)^7(r+4)^4(r+5)\left(\frac{1}{2.4}\right)^{729r}}{\int_0^\infty r^2 e^{-25r} r^{145}(r+1)^{55}(r+2)^{19}(r+3)^7(r+4)^4(r+5)\left(\frac{1}{2.4}\right)^{729r} dr}$$

Numerically, we calculate this is 1.328473×10^{-11}

R-code:

```
y <- (40001:10000000)/100000
x <- (1:10000000)/100000
sum(y^2*exp(-25*y)*y^145*(y+1)^55*(y+2)^19*(y+3)^7*(y+4)^4*(y+5)/2.4^(729*y))/sum(x^2*exp(-
```

(c) Calculate the predictive probability that an individual makes 5 claims next year.

The probability that an individual makes 5 claims next year conditional on $R = r$ is $\frac{r(r+1)(r+2)(r+3)(r+4)}{5!} \left(\frac{1.4}{2.4}\right)^5 \left(\frac{1}{2.4}\right)^r$

The predictive probability is therefore

$$\frac{\left(\frac{1.4}{2.4}\right)^5 \int_0^\infty r^2 e^{-25r} r^{146}(r+1)^{56}(r+2)^{20}(r+3)^8(r+4)^5(r+5)\left(\frac{1}{2.4}\right)^{730r} dr}{5! \int_0^\infty r^2 e^{-25r} r^{145}(r+1)^{55}(r+2)^{19}(r+3)^7(r+4)^4(r+5)\left(\frac{1}{2.4}\right)^{729r} dr}$$

We calculate this numerically:

R-code:

```
sum(x^2*exp(-25*x)*x^146*(x+1)^56*(x+2)^20*(x+3)^8*(x+4)^5*(x+5)/2.4^(729*x))/sum(x^2*exp(-
```

This gives the answer as

$$9.39272 \times \frac{\left(\frac{1.4}{2.4}\right)^5}{5!} = 0.005286815$$

2. An insurance company models loss sizes as following a Pareto distribution with $\alpha = 3$, and finds that the posterior distribution for Θ is a Gamma distribution with $\alpha = 4$ and $\theta = 1400$. Calculate the Bayes estimate for Θ based on a loss function:

(a) $l(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^2$

The value which minimises this expected loss function is the posterior mean of Θ , which is $4 \times 1400 = 5600$.

(b) $l(\hat{\theta}, \theta) = (\hat{\theta} - \theta)^4$

The expected value of this loss function is

$$\begin{aligned}
\int_0^\infty \pi_{\Theta|X}(\theta)(\hat{\theta} - \theta)^4 d\theta &= \int_0^\infty \frac{\theta^3}{1400\Gamma(4)} e^{-\frac{\theta}{1400}} (\hat{\theta} - \theta)^4 d\theta \\
&= \frac{1}{1400\Gamma(4)} \int_0^\infty (\theta^7 - 4\theta^6\hat{\theta} + 6\theta^5\hat{\theta}^2 - 4\theta^4\hat{\theta}^3 + \theta^3\hat{\theta}^4) e^{-\frac{\theta}{1400}} d\theta \\
&= 840 \times 1400^4 - 4 \times 120 \times 1400^3\hat{\theta} + 6 \times 20 \times 1400^2\hat{\theta}^2 - 4 \times 4 \times 1400\hat{\theta}^3 + \hat{\theta}^4
\end{aligned}$$

Numerically, we find that this is minimised by $\hat{\theta} = 6502.119$.

3. An insurance company models claim amounts as following an exponential distribution with mean Θ , where the prior distribution for Θ is a Gamma distribution with $\alpha = 701$ and $\theta = 600$. They observe 700 claims, with mean claim amount \$3,742. Calculate a 95% credibility interval for Θ .

The likelihood of the data is

$$\Theta^{-700} e^{-\frac{2619400}{\Theta}}$$

The prior distribution of Θ is $\pi(\theta) = \frac{\theta^{700} e^{-\frac{\theta}{600}}}{600^{701} \Gamma(701)}$. The posterior distribution is therefore given by

$$\pi_{\Theta|X}(\theta|X) = \frac{\theta^{700} e^{-\frac{\theta}{600}} \theta^{-700} e^{-\frac{2619400}{\theta}}}{\int_0^\infty \theta^{700} e^{-\frac{\theta}{600}} \theta^{-700} e^{-\frac{2619400}{\theta}} d\theta} = \frac{e^{-\frac{\theta}{600} - \frac{2619400}{\theta}}}{\int_0^\infty e^{-\frac{\theta}{600} - \frac{2619400}{\theta}} d\theta}.$$

(a) Using an HPD interval.

The HPD interval is the interval between points θ_1 and θ_2 with equal posterior density. That is

$$\begin{aligned}
e^{-\frac{\theta_1}{600} - \frac{2619400}{\theta_1}} &= e^{-\frac{\theta_2}{600} - \frac{2619400}{\theta_2}} \\
-\frac{\theta_1}{600} - \frac{2619400}{\theta_1} &= -\frac{\theta_2}{600} - \frac{2619400}{\theta_2} \\
\theta_1^2 \theta_2 + 2619400 \times 600 \theta_2 &= \theta_2^2 \theta_1 + 2619400 \times 600 \theta_1 \\
(\theta_1 - \theta_2)(\theta_1 \theta_2 - 2619400 \times 600) &= 0
\end{aligned}$$

Which means either $\theta_1 = \theta_2$ or $\theta_1 \theta_2 = 1571640000$. We want to find the solution to $\theta_1 \theta_2 = 1571640000$ that has a 95% confidence interval.

Substituting $z = \frac{\theta}{600} + \frac{2619400}{\theta}$, we see that z is minimised by $\theta = \sqrt{1571640000}$, and takes the value $z_0 = 2\sqrt{\frac{2619400}{600}}$.

We also see that $\theta = 300z - \sqrt{90000z^2 - 1571640000}$.

Also, we obtain $\frac{dz}{d\theta} = \frac{1}{600} - \frac{2619400}{\theta^2} = \frac{1}{600} - \frac{2619400}{(300z - \sqrt{90000z^2 - 1571640000})^2} = \frac{1}{600} - \frac{2619400}{180000z^2 - 1571640000 - 600z\sqrt{90000z^2 - 1571640000}}$.

Letting $\theta_1 < \theta_2$ satisfy $\theta_1 \theta_2 =$, we find

$$\frac{dz}{d\theta}\Big|_{\theta_2} - \frac{dz}{d\theta}\Big|_{\theta_1} = \frac{2619400}{180000z^2 - 1571640000 - 600z\sqrt{90000z^2 - 1571640000}} - \frac{2619400}{180000z^2 - 1571640000 + 600z\sqrt{90000z^2 - 1571640000}}$$

We are therefore looking to find the value t such that

$$\frac{\int_{z_0}^t \left(\frac{2619400}{180000z^2 - 1571640000 - 600z\sqrt{90000z^2 - 1571640000}} - \frac{2619400}{180000z^2 - 1571640000 + 600z\sqrt{90000z^2 - 1571640000}} \right) e^{-z} dz}{\int_{z_0}^{\infty} \left(\frac{2619400}{180000z^2 - 1571640000 - 600z\sqrt{90000z^2 - 1571640000}} - \frac{2619400}{180000z^2 - 1571640000 + 600z\sqrt{90000z^2 - 1571640000}} \right) e^{-z} dz} = 0.95$$

Numerically, this is solved by $t = 136.090$. The values θ_1 and θ_2 are then the solutions to $\theta = 300 \times 136.090 \pm \sqrt{90000 \times 136.090^2 - 1571640000} = [31069.75, 50584.25]$.

(b) *With equal probability above and below the interval.*

With equal probability above and below the interval, the interval is between the 2.5th percentile and the 97.5th percentile of the posterior distribution. That is, we find the solutions to

$$\frac{\int_0^t e^{-\frac{\theta}{600} - \frac{2619400}{\theta}} d\theta}{\int_0^{\infty} e^{-\frac{\theta}{600} - \frac{2619400}{\theta}} d\theta} = 0.025$$

$$\frac{\int_0^t e^{-\frac{\theta}{600} - \frac{2619400}{\theta}} d\theta}{\int_0^{\infty} e^{-\frac{\theta}{600} - \frac{2619400}{\theta}} d\theta} = 0.975$$

Numerically, the interval is

$$[33693.97, 47350.64]$$

4. *Calculate a conjugate prior distribution for the variance of a normal distribution with mean 0.*

The log-likelihood of the data for a normal distribution with mean 0 and variance σ^2 is

$$-\sum_{i=1}^n \frac{x_i^2}{2\sigma^2} - n \log(\sigma) = -\frac{\sum_{i=1}^n x_i^2}{2\sigma^2} - n \log(\sigma)$$

If the prior distribution is then $\pi(\theta)$, where $\theta = \sigma^2$, then the posterior distribution is proportional to $\pi(\theta)\theta^{-\frac{n}{2}} e^{-\frac{\sum_{i=1}^n x_i^2}{2\theta}}$.

We see that an inverse gamma distribution is therefore a conjugate prior.

Standard Questions

5. An insurance company models number of claims made by an individual in a year as following a Poisson distribution and finds that the posterior distribution for Λ is a Gamma distribution with $\alpha = 4$ and $\theta = 0.02$. The company decides to use an estimate $\hat{\lambda}$ such that the probability of 3 or more claims using $\hat{\lambda}$ is the same as the probability of 3 or more claims under the predictive distribution. Find the value of this $\hat{\lambda}$.

The predictive distribution is a Gamma mixture of Poisson distributions. This is a negative binomial with $r = \alpha, \beta = \theta$. The probability of 3 or more claims is therefore $1 - \left(\frac{1}{1.02}\right)^4 \left(1 + \binom{4}{1} \left(\frac{0.02}{1.02}\right) + \binom{5}{2} \left(\frac{0.02}{1.02}\right)^2\right) = 0.0001442237$.

For the Poisson distribution, we have that the probability of 3 or more claims is $1 - e^{-\lambda} \left(1 + \lambda + \frac{\lambda^2}{2}\right)$, so we need to solve

$$e^{-\lambda} \left(1 + \lambda + \frac{\lambda^2}{2}\right) = 1 - 0.0001442237$$

Numerically, we find the solution to this is $\lambda = 0.0976423$.

Bonus Question

6. An insurance company models loss amounts as following a Weibull distribution with $\tau = 3$. It uses the inverse gamma prior for the unknown parameter Θ , with parameters $\alpha = 3$ and $\theta = 6000$. This is a conjugate prior, and the posterior distribution after observing N observations X_1, \dots, X_n with $\sum_{i=1}^n X_i^\tau = t$ is inverse gamma with $\alpha = 3 + N$, and $\theta = \frac{1}{\frac{1}{6000} + t}$. Calculate the probability that the posterior probability of $\Theta > 3000$ is more than 0.05, after a sample of 10 observations.

For given values of N and t , the posterior probability that $\Theta > 3000$ is the probability that an inverse gamma distribution with $\alpha = 3 + N$ and $\theta = \frac{1}{\frac{1}{6000} + t}$ is more than 3000. This is the probability that a Gamma distribution with $\alpha = 3 + N$ and $\theta = \frac{1}{6000} + t$ is less than $\frac{1}{3000}$. We have that $N = 10$, so this probability is the quantile of Gamma distribution with $\alpha = 13$ corresponding to the value $\frac{\left(\frac{1}{3000}\right)}{\left(\frac{1}{6000} + t\right)} = \frac{2}{1+6000t}$. That is, the overall probability is

$$\begin{aligned} & \int_0^\infty f_T(t) \int_{\frac{2}{1+6000t}}^\infty \frac{x^{12} e^{-x}}{12!} dx dt \\ &= \int_0^\infty \int_{\frac{1}{3000x} - \frac{1}{6000}}^\infty \frac{x^{12} e^{-x}}{12!} f_T(t) dt dx \end{aligned}$$

for each X_i , we have that X_i^τ follows an exponential distribution, with mean Θ . This means that T follows a gamma distribution with parameters 10 and Θ . This gives a probability of

$$\begin{aligned}
& \int_0^\infty f_T(t) \int_{\frac{2}{1+6000t}}^\infty \frac{x^{12} e^{-x}}{12!} dx dt \\
&= \int_0^\infty \int_{\frac{1}{3000x} - \frac{1}{6000}}^\infty \frac{x^{12} e^{-x}}{12!} \frac{t^9 e^{-\frac{t}{\theta}}}{\theta^{10} 10!} dt dx \\
&= \int_0^\infty \int_{\frac{1}{3000x\theta} - \frac{1}{6000\theta}}^\infty \frac{x^{12} e^{-x}}{12!} \frac{u^9 e^{-u}}{10!} du dx \\
&= \int_0^\infty \int_{\frac{1}{3000x\theta} - \frac{1}{6000\theta}}^\infty \frac{x^{12} u^9 e^{-(x+u)}}{12! \times 10!} du dx
\end{aligned}$$

Substituting $a = x + u$, this becomes

$$\begin{aligned}
& \int_l^\infty \frac{e^{-a}}{12! \times 10!} \int_{h(a)}^a (a-u)^{12} u^9 du da \\
&= \int_l^\infty \frac{e^{-a}}{12! \times 10!} \int_0^{k(a)} v^{12} (a-v)^9 dv da \\
&= \int_l^\infty \frac{e^{-a}}{12! \times 10!} \left(\frac{a^9 k(a)^{13}}{13} - 9 \frac{a^8 k(a)^{14}}{14} + \dots - \frac{k(a)^{22}}{22} \right) da
\end{aligned}$$

where $l = \frac{2}{\sqrt{3000\theta}} - \frac{1}{6000\theta}$ and $k(a) = a - \frac{(a-c) + \sqrt{(a-c)^2 + 4(ac-2c)}}{2} = \frac{(a+c) - \sqrt{(a+c)^2 - 8c}}{2}$ where $c = \frac{1}{6000\theta}$